# Housing Price Forecasting based on Stochastic Time Series Model

## Jianhua GUO

Department of Mathematics, Hunan University of Science and Technology, Xiangtan, China jhguo888@163.com

#### Abstract:

Whether the real estate industry robustly develops or not largely affects national macroeconomic development and national quality of life. By analyzing monthly averaged prices of commercial residential building in Changsha City from Jan, 2002 to Dec, 2011, this paper aims to reasonably construct a forecasting model to predict short-term housing price trend and affords reference to homebuyer and investors, what's more, affords technical support to government's policy making. First, how to select rational forecasting model is discussed, and then a price forecasting ARMA model is constructed, lastly, empirical analysis is put forward.

Keywords: housing price, forecasting, stochastic time series, ARMA model

## **1. INTRODUCTION**

Commercial housing price affects national macroeconomic development and national quality of life in a large extent. Variation of commercial housing price also deeply impact the running of national economy, so, study on housing price is a significant job.

So far, many scholars studied influencing factor of housing price and acquired plenty of achievement, but the quantitative research about housing price is in the infant stage. Among which, Malpezzi(1999) studied 133 American cities' commercial housing price index and concluded that the housing price is not a random walk but partly predictable. Clapp J.H. and Giaccotto C.(2002), by means of autoregression model and relying on the past quarterly housing prices during 1976 and 1997, forecasted the next quarterly price. Seko(2003) studied the relationship between dwelling price and macro economy in Japan. Wedyawati W. and Meiliu L.(2004) made a price for commercial residential building by constructing data mining system between homebuyer and seller and by aid of regression technique. Okmyung B.(2004) compared the parametric and nonparametric price forecasting methods. Bradford C. et al(2004) compared 4 housing price forecasting methods including Dubinrsquos -krigining local forecasting model, Clapprsquou local regression model, local instance forecasting and nearest-neighbor method. Anglin(2006), by introducing the averaged growth rate of lagged 3-term housing prices, CPI, interest rate of mortgage lending and rate of unemployment and constructing VAR model, forecasted the housing price variation in Toronto. Liu, J.G. and Zhang, X.L. et al (2006) estimated property price making use of fuzzy-neural network technique. Zhou, Wei-Xing and Didier, S. (2008) analyzed the housing price indexes of Las Vegas during June, 1983 and March, 2005 and found out the seasonal trend, then, they forecasted the next year's housing price index by moving average method. Min Hwang John M. Quigley(2010), based on Singaporean housing prices during 1990 and 2000, constructed a price forecasting model and found that the housing price submitted to mean-reversion model and was in connection with region.

## 2. MARKET MODELS

Time series analysis is a widely used quantitative analyzing method and time series model, including linear model and nonlinear model, is an important kind of economic model to depicting economic data's evolving process. Linear model comprise of AR model, MA model and ARMA model, etc., as to any time

series, we can ascertain whether it is stationary time series or not by unit-root testing, i.e., the time series is stationary when all characteristic roots are bigger than 1.what's more, if the autocorrelation coefficients with order bigger than 3 all located in confidence interval and converging to zero, we also consider the time series to be stationary, and vice verse.

#### 2.1 AR model

As to time series  $\{X_t\}$ , if there is

$$X_{t} = \phi_{1}X_{t-1} + \phi_{2}X_{t-2} + \dots + \phi_{p}X_{t-p} + \varepsilon_{t}$$
(1)

Then,  $\{X_t\}$  is called an autoregression time series, and (1) is called an AR(p) model. Where  $\phi_1, \phi_2, \dots, \phi_p$  are autoregression coefficients,  $\varepsilon_t \sim N(0, \sigma^2)$  and  $E(\varepsilon_t X_{t-i}) = 0, i = 1, \dots, p$ . If lagged operator B is introduced, then,  $X_{t-1} = BX_t$ ,  $X_{t-2} = B^2 X_t$ , ....,  $X_{t-p} = B^p X_t$ , and (1) can be transferred into

$$X_{t} = \phi_{1}BX_{t} + \phi_{2}B^{2}X_{t} + \dots + \phi_{p}B^{p}X_{t} + \varepsilon_{t}$$
(2)

Further more, denoting  $\phi(B) = 1 - \phi_1 B - \phi_2 B^2 - \dots - \phi_p B^p$ , then

$$\phi(B)X_t = \mathcal{E}_t \tag{3}$$

We call  $\phi(B) = 0$  is the characteristic equation of the autoregression model AR(p), which is called a stationary autoregression model, if all of characteristic roots  $\lambda_i$ ,  $i = 1, \dots, p$  of  $\phi(B) = 0$  are located outside the unit circle, i.e.,

$$|\lambda_i| > 1, i = 1, \cdots, p \tag{4}$$

A time series evolving as a stationary model AR(p) is called a stationary AR(p) time series. The necessary and sufficient conditions for a stationary time series to be an AR(p) time series is that the partial correlation function is p-step clipped, i.e.,  $\{X_t\}$  only has p nonzero partial correlation coefficients  $\psi_{kk}$ .

In fact, because of the randomicity of sample, all  $\psi_{kk}$  will not equal to zero though k>p, instead,  $\psi_{kk} \sim N(0,1/n)$ , where n denotes the sample length, and as for enough large n, there is  $P\{|\psi_{kk}| < 2/\sqrt{n}\} = 0.955$ , thus, we generally recur to interval-testing method to ascertain whether a time series is an AR(p) time series or not.

Note, if  $\{X_t\}$  is an AR(p) time series, the partial correlation function  $\psi_{kk}$  is p-step clipped, but its autocorrelation function  $\rho_k$  is trailing, i.e., attenuating according to a negative exponential function.

#### 2.2 MA model

As to time series  $\{X_t\}$ , if

$$X_{t} = \varepsilon_{t} - \theta_{1}\varepsilon_{t-1} - \theta_{2}\varepsilon_{t-2} - \dots - \theta_{q}\varepsilon_{t-q}$$
(5)

Then,  $\{X_t\}$  is called a moving average time series, and (5) is called a MA(q) model. Where  $\theta_1, \theta_2, \dots, \theta_p$  are moving average coefficients. Similarly, if lagged operator B is introduced, and denoting  $\theta(B) = 1 - \theta_1 B - \theta_2 B^2 - \dots - \theta_q B^q$ , (5) can be transferred into

$$X_t = \theta(B)\mathcal{E}_t \tag{6}$$

We call  $\theta(B) = 0$  is the characteristic equation of the moving average model MA(q), which is called a reversible moving average model, if all of characteristic roots  $\lambda_i$ ,  $i = 1, \dots, p$  of  $\theta(B) = 0$  are located outside the unit circle, i.e.,  $|\lambda_i| > 1, i = 1, \dots, p$ 

A time series evolving as a reversible model MA(q) is called a reversible MA(q) time series. The necessary and sufficient conditions for a stationary time series to be an MA(q) time series is that the auto correlation function is q-step clipped, i.e.,  $\{X_t\}$  only has q nonzero auto correlation coefficients  $\rho_k$ . It means that, as to a MA(q) time series,  $X_t$  and  $X_s$  are irrelevant when |t - s| > q.

In fact, because of the randomicity of sample, all  $\rho_k$  will not equal to zero though k>q, instead,

$$\rho_k \sim N(0, \frac{1}{n}(1+2\sum_{i=1}^q \rho_i^2))$$
, where n denotes the sample length, and as for enough large n, there is

 $P\{|\rho_k| < \frac{2}{\sqrt{n}}\sqrt{1+2\sum_{i=1}^q \rho_i^2}\} = 0.955$ , thus, we generally recur to interval-testing method to ascertain

whether a time series is an MA(q) time series or not.

Similarly as AR(p) time series, if  $\{X_t\}$  is a MA(q) time series, the auto correlation function  $\rho_k$  is q-step

clipped, but its partial correlation function  $\psi_{kk}$  is trailing, i.e., attenuating according to a negative exponential function.

## 2.3 ARMA model

If time series  $\{X_t\}$  satisfies

$$X_{t} = \phi_{1}X_{t-1} + \phi_{2}X_{t-2} + \dots + \phi_{p}X_{t-p} + \varepsilon_{t} - \theta_{1}\varepsilon_{t-1} - \theta_{2}\varepsilon_{t-2} - \dots - \theta_{q}\varepsilon_{t-q}$$
(7)

Then, we call a autoregression-moving average time series, and call (7) a autoregression-moving average model, i.e., ARMA(p,q) model.

Where  $\phi_1, \phi_2, \dots, \phi_p$ ,  $\theta_1, \theta_2, \dots, \theta_p$  and  $\varepsilon_{t-1}, \dots, \varepsilon_{t-q}$  are same as in subsection 2.1 and subsection 2.2.

Also, when introducing lagged operator B, (7) become into

$$\phi(B)X_t = \theta(B)\varepsilon_t \tag{8}$$

Obviously, AR(p) model and MA(q) model are special cases of ARMA(p,q) model.

If a stationary time series  $\{X_t\}$  has clipped partial correlation function and trailing autocorrelation function,

 $\{X_t\}$  is an AR time series; If a stationary time series  $\{X_t\}$  has clipped autocorrelation function and trailing

partial correlation function,  $\{X_{t}\}$  is a MA time series; If the partial correlation function and autocorrelation

function of  $\{X_t\}$  all are trailing, it is an ARMA time series.

#### Note

1) In reality, many time series are not stationary time series, then, we can translate them into stationary ones by difference method.

2) We can make use of the residual error testing method or fitting testing method to decide whether an AR(p) model, MA(q) model or an ARMA(p,q) model is appropriate to depict a time series or not.

#### 3. SAMPLING AND MODELING

#### 3.1 Sampling and stationarity test

In this paper, we select monthly housing prices of Changsha City from Jan,2009 to Dec,2011(shown in figure 1) as within sample data for modeling, while selecting monthly housing prices of January and February, 2012 to detect price forecasting precision.

First, as to figure 1, we can see that there is visible uptrend component in price series and the price series is non-periodic, so, it is reasonable to decide the price series be non-stationary. In addition, the autocorrelation function is attenuating as a smooth exponential function while the partial correlation function is clipped with positive peak value, which, just as unit-root testing, shows the price series is non-stationary, see figure 2 and figure 3, respectively. Therefore, we make differential operation for the original housing price series (see figure 4), according to its autocorrelation function and partial correlation function, as well as unit-root testing (see figure 5 and figure 6, respetively), we can ascertain that the differenced series is stationary.



Figure 1 The housing price series of Changsha City during Jan, 2009 and Dec, 2011 (Unit: CNY per square meter)

Date: 03/12/12 Time: 22:58 Sample: 2009M01 2011M12 Included observations: 36

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
		1	0.919	0.919	32.990	0.000
1	1 1	2	0.844	0.001	61.663	0.000
1	111	3	0.777	0.012	86.722	0.000
		4	0.700	-0.105	107.65	0.000
I	101	5	0.618	-0.071	124.52	0.000
		6	0.530	-0.105	137.32	0.000
	וםי	7	0.439	-0.074	146.43	0.000
· 🗖		8	0.352	-0.045	152.50	0.000
· 🗖 ·		9	0.272	-0.012	156.25	0.000
· 🗖 ·		10	0.193	-0.048	158.21	0.000
1 <b>p</b> 1	יםי	11	0.110	-0.085	158.86	0.000
111		12	0.033	-0.037	158.92	0.000
1 🛛 1		13	-0.029	0.022	158.97	0.000
יםי	I <b>(</b> I	14	-0.089	-0.048	159.47	0.000
		15	-0.140	0.002	160.74	0.000
		16	-0.176	0.028	162.87	0.000
		17	-0.242	-0.245	167.10	0.000
		18	-0.290	0.011	173.49	0.000
		19	-0.318	0.035	181.60	0.000
	יוי	20	-0.335	0.045	191.18	0.000
		21	-0.372	-0.187	203.81	0.000
		22	-0.398	0.002	219.28	0.000
		23	-0.414	-0.028	237.32	0.000
		24	-0.424	-0.008	257.79	0.000
		25	-0.421	0.014	279.80	0.000
	יוי	26	-0.408	0.043	302.59	0.000
		27	-0.393	0.007	326.04	0.000
		28	-0.371	-0.014	349.62	0.000
		29	-0.342	-0.023	372.44	0.000
	יו	30	-0.305	0.045	393.62	0.000
		31	-0.261	0.045	412.24	0.000
' <b>-</b> '		32	-0.215	-0.014	428.05	0.000
' 🗖 '	וו	33	-0.166	0.062	440.58	0.000
' <b>□</b> '		34	-0.118	-0.022	450.09	0.000
1 <b>(</b> )		35	-0.062	0.030	455.35	0.000

Figure 2 Figure of autocorrelation function and partial correlation function of price series

Null Hypothesis: SER01 has a unit root Exogenous: Constant, Linear Trend Lag Length: 2 (Automatic based on AIC, MAXLAG=9)

		t-Statistic	Prob.*
Augmented Dickey-Fu	ller test statistic	-0.105300	0.9926
Test critical values:	1% level	-4.262735	
	5% level	-3.552973	
	10% level	-3.209642	

\*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation Dependent Variable: D(SER01) Method: Least Squares Date: 03/12/12 Time: 23:00 Sample (adjusted): 2009M04 2011M12 Included observations: 33 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
SER01(-1) D(SER01(-1)) D(SER01(-2)) C @TREND(2009M01)	-0.015266 -0.343628 -0.308079 239.9401 -2.441900	0.144974 0.209131 0.193387 520.4076 11.93560	-0.105300 -1.643121 -1.593068 0.461062 -0.204590	0.9169 0.1115 0.1224 0.6483 0.8394
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.192827 0.077516 114.1634 364931.8 -200.4558 1.672240 0.184368	Mean depend S.D. depende Akaike info cr Schwarz crite Hannan-Quin Durbin-Watsc	lent var int var iterion rion n criter. on stat	66.18182 118.8632 12.45187 12.67861 12.52816 1.950402

#### Figure 3 Unit-root test results



Figure 4 The differenced price series of the original housing price series

Date: 03/12/12 Time: 23:10 Sample: 2009M02 2011M12 Included observations: 35

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Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
		1 -0.205	-0.205	1.6010	0.206
		2 -0.178	-0.230	2.8493	0.241
ı 🗖 i		3 0.185	0.102	4.2294	0.238
1 🚺 1		4 -0.027	-0.000	4.2601	0.372
1 <b>j</b> 1	<u> </u>	5 0.027	0.085	4.2908	0.508
ı 🗖 i	I 🗖 I	6 0.190	0.212	5.8970	0.435
	<u> </u>	7 -0.058	0.062	6.0532	0.534
		8 -0.165	6-0.132	7.3567	0.499
		9 -0.018	-0.177	7.3722	0.598
1 <b>p</b> 1	1 🚺 1	10 0.082	-0.036	7.7163	0.657
1 <b>D</b> 1		11 0.114	0.145	8.4167	0.676
I 🖬 I		12 -0.067	0.031	8.6664	0.731
1 <b>j</b> 1	וםי	13 0.028	0.112	8.7126	0.794
I 🛛 I		14 -0.082	-0.045	9.1247	0.823
1 1		15 0.003	-0.014	9.1254	0.871
1 <b>p</b> 1		16 0.085	-0.047	9.6180	0.886
1 1		17 0.009	-0.037	9.6244	0.919
יםי		18 -0.133	-0.118	10.972	0.896
		19 -0.017	-0.038	10.995	0.924
1 🛛 1		20 -0.045	6 -0.064	11.169	0.942
1 🛛 1	I [ I	21 -0.036	i -0.053	11.292	0.957
		22 -0.038	-0.132	11.437	0.968
I 🛛 I		23 -0.055	i -0.099	11.763	0.974
1 🛛 1		24 -0.037	-0.042	11.927	0.981
	וןי	25 -0.010	0.036	11.939	0.987
1 1		26 -0.005	0.008	11.943	0.991
		27 -0.010	0.005	11.958	0.994
		28 0.009	0.017	11.973	0.996
	וםי	29 0.009	0.048	11.992	0.998
	ווי	30 0.021	0.031	12.105	0.998
		31 -0.013	0.004	12.163	0.999
		32 -0.020	-0.037	12.336	0.999
		33 -0.027	-0.033	12.803	0.999
1 1 1		34 -0.016	-0.045	13.119	1.000

Figure 5 Figure of autocorrelation and partial correlation functions of the differenced price series

Null Hypothesis: DPRICE has a unit root Exogenous: Constant, Linear Trend Lag Length: 1 (Automatic based on AIC, MAXLAG=8)

		t-Statistic	Prob.*
Augmented Dickey-Fu	ller test statistic	-6.006053	0.0001
Test critical values:	1% level	-4.262735	
	5% level	-3.552973	
	10% level	-3.209642	

\*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation Dependent Variable: D(DPRICE) Method: Least Squares Date: 03/12/12 Time: 23:11 Sample (adjusted): 2009M04 2011M12 Included observations: 33 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
DPRICE(-1) D(DPRICE(-1)) C @TREND(2009M02)	-1.671236 0.316093 181.7769 -3.677396	0.278259 0.174724 52.74650 2.151592	-6.006053 1.809097 3.446237 -1.709151	0.0000 0.0808 0.0018 0.0981
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.673774 0.640026 112.2000 365076.3 -200.4623 19.96512 0.000000	Mean depend S.D. depende Akaike info cri Schwarz crite Hannan-Quin Durbin-Watsc	lent var nt var terion rion n criter. ın stat	-5.333333 187.0068 12.39166 12.57305 12.45269 1.956295

Figure 6 Unit-root test results of the differenced price series

## 3.2 Modeling and Forecasting

According to figure 5, all autocorrelation coefficients and partial correlation coefficients locate in confidence interval when k>1, and convergence to zero, so, we can model the original price series with one of the ARI(1,1),MAI(1,1) or ARIMA(1,1,1) models. Relying on the EVIEWS software, the AIC values of three models are 12.44865, 12.38843, 12.44639, respectively, and by means of theory of minimizing criterion function, we first select MAI(1,1) as our candidate time series model. However, the AIC value is not the sufficient condition for selecting the optimal model. Here, we do significance test for parameters of the selected model or do test for randomness of residual error, if pass the test, the selected model with least AIC value is the optimal model, or else, select model with minor AIC value and decide whether parametric test or residual error test passed, and so on, until the appropriate model is decided.

In this paper, after trial and error, the ARIMA(1,1,1) is identified as the optimal one. Come next, we estimate model's parameters. Among all favorite parameter estimation methods, including moment estimation, maximum likelihood estimation and nonlinear least-square(NLS) estimation, etc, here, in EVIEWS software, we adopt NLS method to estimate parameters, results as shown in figure 7.

Dependent Variable: DPRICE Method: Least Squares Date: 03/13/12 Time: 17:24 Sample (adjusted): 2009M03 2011M12 Included observations: 34 after adjustments Convergence achieved after 16 iterations MA Backcast: 2009M02

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C AR(1) MA(1)	69.06684 0.112233 -0.395474	14.00293 0.535655 0.510031	4.932315 0.209525 -0.775392	0.0000 0.8354 0.4440
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.081758 0.022516 117.0038 424386.8 -208.5886 1.380074 0.266587	Mean dependent var S.D. dependent var Akaike info criterion Schwarz criterion Hannan-Quinn criter. Durbin-Watson stat		69.17647 118.3437 12.44639 12.58107 12.49232 1.954560
Inverted AR Roots Inverted MA Roots	.11 .40			

Figure 7 estimation of ARMA model

Thus, the expressions of 1-order difference of original housing price series is as following

 $\Delta price_{t} = 69.0668 + 0.11223 \Delta price_{t-1} + \varepsilon_{t} - 0.395474 \varepsilon_{t-1}$ (9)

The randomicity of residual error series, i.e., all autocorrelation coefficients come close to zero when K>1, is an important indicator to decide a model be appropriate or not. Shown as in figure 8, all residual errors are minute and autocorrelation coefficients and partial correlation coefficients all locate in confidence interval, as well as, the P-values of Q statistic are much bigger than 0.05, so, it's reasonable to consider the residual error series as a white noise series and the ARIMA(1,1,1) model is appropriate to depict the housing price series.

Date: 03/13/12 Time Sample: 2009M02 20 Included observation	e: 17:57 011M12 s: 34					
Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
		1 1 2 3 4 4 5 6 7 7 8 9 10 11 12 13 14 5 16 17 18 19 20 21 22 23 24 25 26 27 7 28 9 30 31 2	-0.228 -0.168 0.203 -0.009 0.021 0.187 -0.066 -0.165 -0.017 0.049 0.079 -0.066 0.010 -0.010 -0.061 -0.011 -0.044 -0.043 -0.044 -0.044 -0.051 -0.051 -0.005 -0.005 -0.007 -0.011 0.011 0.011 0.012 0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.026 -0.021 -0.025 -0.02	-0.228 -0.232 0.114 0.039 0.223 0.057 -0.150 -0.218 -0.114 0.033 0.126 -0.040 0.040 -0.033 0.126 -0.040 -0.051 -0.074 -0.074 -0.036 -0.036 -0.036 0.013 -0.045 -0.033 0.025 0.033 0.025 -0.015 -0.033	1.9276 3.0124 4.6334 4.6367 4.6548 6.1913 6.3915 7.6712 7.6852 7.8060 8.1374 8.3741 8.3794 9.0682 9.3748 8.3741 8.3794 9.0682 9.3748 9.3748 9.3748 9.3933 9.4019 10.240 10.240 10.240 11.187 11.200 11.234 11.277 11.482	0.165 0.222 0.201 0.327 0.459 0.460 0.566 0.566 0.566 0.701 0.755 0.818 0.827 0.856 0.927 0.924 0.940 0.940 0.970 0.940 0.970 0.960 0.970 0.988 0.992 0.995 0.995 0.9999 0.9999
1 1 1		33	-0.013	-0.033	11.705	1.000

Figure 8 white noise test of the residual error series

According to the constructed ARIMA(1,1,1) model, we forecast the next two months' housing prices in table 1 as follows

Month	Actual housing price	Forecasted value	Relative error			
Jan,2012	6154	6202	0.78%			
Feb,2012	6129	6225	1.56%			

Table 1 the forecasted housing price and relative error to actual price

#### 4. CONCLUSION

The ARMA model forecasts next values based on preceding time series data and it can be used to forecast short-term housing price. The empirical analysis indicates that this model's forecasting precision is satisfactory, and only with some simple modification, this model may be used in other industries. Generally speaking, housing price affected by many factors is nonlinear changing, but there few of literatures study on housing price series from nonlinear point of view, so, we can say that the ARMA model is a comparatively excellent forecasting model for its convenience of application and its strict mathematical assurance, besides, differ from other models such as regression model relying on great deal of related data , the ARMA model only rely on historical data of some certain variable, which make its unique predominance in short-term forecast.

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